

# Course 02402 Introduction to Statistics

## Lecture 10: Inference for proportions

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## Overview

- 1 Introduction
- 2 Confidence interval for one proportion
  - Sample size determination (planning)
- 3 Hypothesis test for one proportion
- 4 Confidence interval and hypothesis test for two proportions
- 5 Hypothesis test for several proportions
- 6 Analysis of contingency tables

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## Different analysis/data situations in 02402

### Mean of quantitative data

- Hypothesis test/CI for one mean (one sample)
- Hypothesis test/CI for two means (two samples)
- Hypothesis test/CI for several means ( $K$  samples)

### Today: Proportions

- Hypothesis test/CI for one proportion
- Hypothesis test/CI for two proportions
- Hypothesis test for several proportions
- Hypothesis test for several “multi-categorical” proportions

## Estimation of proportions

- Estimation of a proportion/probability, by observing how many times  $x$  an event has occurred in  $n$  (independent) trials:

$$\hat{p} = \frac{x}{n}$$

- Note that  $\hat{p} \in [0; 1]$ .
- Example: A dice is thrown  $n = 100$  times. In  $x = 20$  cases the outcome was 3. Then,  $\hat{p}$  is the estimated probability of throwing a 3.

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## Confidence interval for one proportion

### Method 7.3

If we have a **large sample**, then a  $(1 - \alpha)100\%$  confidence interval for  $p$  is:

$$\frac{x}{n} - z_{1-\alpha/2} \sqrt{\frac{\frac{x}{n}(1-\frac{x}{n})}{n}} < p < \frac{x}{n} + z_{1-\alpha/2} \sqrt{\frac{\frac{x}{n}(1-\frac{x}{n})}{n}}$$

### How?

Follows from **approximating** the binomial distribution by the normal distribution.

### A rule of thumb

Suppose that  $X \sim \text{binom}(n, p)$ . The normal distribution is a good approximation of the binomial distribution if  $np$  and  $n(1-p)$  (expected no. of successes and failures, respectively) are both greater than 15.

## Confidence interval for one proportion

### Mean and variance of binomial distribution, Chapter 2.21

$$\begin{aligned} E(X) &= np \\ \text{Var}(X) &= np(1-p) \end{aligned}$$

### This means that

$$\begin{aligned} E(\hat{p}) &= E\left(\frac{X}{n}\right) = \frac{np}{n} = p \\ \text{Var}(\hat{p}) &= \text{Var}\left(\frac{X}{n}\right) = \frac{1}{n^2} \text{Var}(X) = \frac{p(1-p)}{n} \end{aligned}$$

## Example 1

Left-handedness:

$p$  = Proportion of left-handed people in Denmark

and/or:

Female engineering students:

$p$  = Proportion of female engineering students

## Example 1

Left-handedness:

$$\sqrt{\frac{\hat{p}(1-\hat{p})}{n}} = \sqrt{\frac{10/100(1-10/100)}{100}} = 0.03$$

$$0.10 \pm 1.96 \cdot 0.03 = 0.10 \pm 0.059 = [0.041, 0.159]$$

Better “small sample” method - the “plus 2-approach” (Remark 7.7)

Use the same formula on  $\tilde{x} = 10 + 2 = 12$  and  $\tilde{n} = 100 + 2 + 2 = 104$ :

$$\sqrt{\frac{\tilde{p}(1-\tilde{p})}{\tilde{n}}} = \sqrt{\frac{12/104(1-12/104)}{104}} = 0.031$$

$$0.115 \pm 1.96 \cdot 0.031 = 0.115 \pm 0.061 = [0.054, 0.177]$$

## The Margin of Error (ME)

The Margin of Error

with  $(1 - \alpha)100\%$  confidence becomes:

$$ME = z_{1-\alpha/2} \sqrt{\frac{p(1-p)}{n}}$$

where  $p$  may be estimated using  $\hat{p} = \frac{x}{n}$ .

The Margin of Error:

- Corresponds to half the width of the  $(1 - \alpha)100\%$  confidence interval.
- Describes the “minimum desired precision” of the estimate  $\hat{p}$ .

## Sample size determination

Design of experiments:

How large should the sample size  $n$  be in order to obtain the desired precision?

Method 7.13

If you want a Margin of Error, ME, with  $(1 - \alpha)100\%$  confidence, then you need the following sample size:

$$n = p(1-p) \left( \frac{z_{1-\alpha/2}}{ME} \right)^2$$

## Sample size determination

### Method 7.13

If you want a Margin of Error, ME, with  $(1 - \alpha)100\%$  confidence, and you do *not* have a reasonable guess of  $p$ , then you need the following sample size:

$$n = \frac{1}{4} \left( \frac{z_{1-\alpha/2}}{\text{ME}} \right)^2$$

since the worst case approach is given by:  $p = \frac{1}{2}$

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## Example 1 - continued

### Left-handedness:

Suppose that we want  $\text{ME} = 0.01$  (with  $\alpha = 0.05$ ) – what should  $n$  be?

Assume  $p \approx 0.10$ :

$$n = 0.1 \cdot 0.9 \left( \frac{1.96}{0.01} \right)^2 = 3467.4 \approx 3468$$

Without any assumption on the size of  $p$ :

$$n = \frac{1}{4} \left( \frac{1.96}{0.01} \right)^2 = 9604$$

## Steps of a hypothesis test - an overview (repetition)

- 1 Formulate the hypothesis and choose the level of significance  $\alpha$  (i.e. the “risk-level”).
- 2 Use the data to calculate the value of the test statistic.
- 3 Calculate the p-value using the test statistic and the relevant distribution. Compare the  $p$ -value to the significance level  $\alpha$  and draw a conclusion.
- 4 (Alternatively, draw a conclusion based on the relevant critical value(s)).

## Hypothesis test for one proportion

The null and alternative hypothesis for one proportion  $p$ :

$$H_0: p = p_0$$

$$H_1: p \neq p_0$$

We either accept  $H_0$  or reject  $H_0$ .

## Testing the hypothesis: The test statistic

### Theorem 7.10 and Method 7.11

If the sample size is sufficiently large (if  $np_0 > 15$  and  $n(1 - p_0) > 15$ ), we use the following test statistic:

$$z_{\text{obs}} = \frac{x - np_0}{\sqrt{np_0(1 - p_0)}}$$

Under the null hypothesis, the random variable  $Z$  (approximately) follows a standard normal distribution,  $Z \sim N(0, 1^2)$ .

## Testing the hypothesis: $p$ -value and conclusion (Method 7.11)

Find the  $p$ -value (evidence against the null hypothesis):

- $2P(Z > |z_{\text{obs}}|)$

Test using the critical value:

Reject null hypothesis if  $z_{\text{obs}} < -z_{1-\alpha/2}$  or  $z_{\text{obs}} > z_{1-\alpha/2}$ .

## Example 1 - continued

Is half of the Danish population left handed?

$$H_0: p = 0.5, H_1: p \neq 0.5$$

Test statistic:

$$z_{\text{obs}} = \frac{x - np_0}{\sqrt{np_0(1 - p_0)}} = \frac{10 - 100 \cdot 0.5}{\sqrt{100 \cdot 0.5(1 - 0.5)}} = -8$$

$p$ -value:

$$2 \cdot P(Z > 8) = 1.2 \cdot 10^{-15}$$

There is very strong evidence against the null hypothesis - we reject it (with  $\alpha = 0.05$ ).

## Example 1 - continued

Testing the hypothesis in R

```
prop.test(10, 100, p = 0.5, correct = FALSE)

##
## 1-sample proportions test without continuity correction
##
## data: 10 out of 100, null probability 0.5
## X-squared = 64, df = 1, p-value = 1e-15
## alternative hypothesis: true p is not equal to 0.5
## 95 percent confidence interval:
##  0.05523 0.17437
## sample estimates:
##      p
## 0.1
```

## Overview

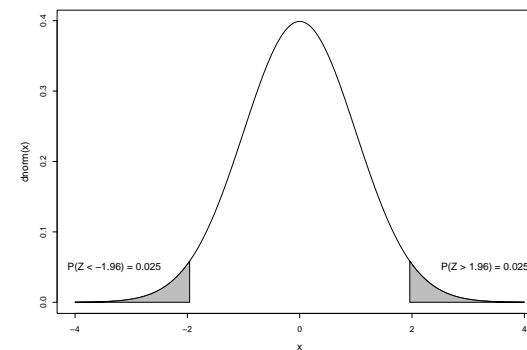
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## Example 1 - continued

Using the critical value instead:

$$z_{0.975} = 1.96$$

As  $z_{\text{obs}} = -8$  is (much) less than  $-1.96$  we reject the hypothesis.



## Confidence interval for (the difference between) two proportions

### Method 7.15

$$(\hat{p}_1 - \hat{p}_2) \pm z_{1-\alpha/2} \cdot \hat{\sigma}_{\hat{p}_1 - \hat{p}_2}$$

where

$$\hat{\sigma}_{\hat{p}_1 - \hat{p}_2} = \sqrt{\frac{\hat{p}_1(1 - \hat{p}_1)}{n_1} + \frac{\hat{p}_2(1 - \hat{p}_2)}{n_2}}$$

Rule of thumb:

Both  $n_i p_i \geq 10$  and  $n_i(1 - p_i) \geq 10$  for  $i = 1, 2$ .

## Hypothesis test for two proportions, Method 7.18

### Two sample proportions hypothesis test

Comparing two proportions (shown here for a two-sided alternative)

$$H_0: p_1 = p_2$$

$$H_1: p_1 \neq p_2$$

The test statistic:

$$z_{\text{obs}} = \frac{\hat{p}_1 - \hat{p}_2}{\sqrt{\hat{p}(1-\hat{p})\left(\frac{1}{n_1} + \frac{1}{n_2}\right)}}, \quad \text{where } \hat{p} = \frac{x_1 + x_2}{n_1 + n_2}$$

And for large samples:

Use the standard normal distribution again.

## Example 2

Is there a relation between the use of birth control pills and the risk of a blood clot in the heart?

In a study (USA, 1975) the connection between birth control pills and the risk of a blood clot in the heart was investigated.

	Blood clot	No blood clot
B. C. pill	23	34
No B. C. pill	35	132

Carry out a test to check if there is any connection between the use of birth control pills and the risk of a blood clot in the heart. Use a significance level of  $\alpha = 0.05$ .

## Example 2

In a study (USA, 1975) the connection between birth control pills and the risk of blood clot in the heart was investigated.

	Blood clot	No blood clot
B. C. pill	23	34
No B. C. pill	35	132

Estimates within each sample

$$\hat{p}_1 = \frac{23}{57} = 0.4035, \quad \hat{p}_2 = \frac{35}{167} = 0.2096$$

Common estimate:

$$\hat{p} = \frac{23 + 35}{57 + 167} = \frac{58}{224} = 0.2589$$

## Example 2 - continued

`prop.test` for equality of two proportions in R

```
# Read data table into R
pill.study <- matrix(c(23, 34, 35, 132),
                    ncol = 2, byrow = TRUE)
colnames(pill.study) <- c("Blood Clot", "No Clot")
rownames(pill.study) <- c("Pill", "No pill")
pill.study

# Test whether probabilities are equal for the two groups
prop.test(pill.study, correct = FALSE)
```

## Example 2 - continued

prop.test for equality of two proportions in R

```
##           Blood Clot No Clot
## Pill           23      34
## No pill        35     132
```

```
##
## 2-sample test for equality of proportions without continuity
## correction
##
## data: pill.study
## X-squared = 8.3, df = 1, p-value = 0.004
## alternative hypothesis: two.sided
## 95 percent confidence interval:
##  0.05239 0.33546
## sample estimates:
## prop 1 prop 2
## 0.4035 0.2096
```

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## Hypothesis test for several proportions

### The comparison of $c$ proportions

In some cases, we might be interested in determining whether two or more binomial distributions have the same parameter  $p$ . That is, we are interested in testing the null hypothesis:

$$H_0: p_1 = p_2 = \dots = p_c = p$$

vs. the alternative that at least two proportions are different.

## Hypothesis test for several proportions

Table of observed counts for  $c$  samples:

	Sample 1	Sample 2	...	Sample $c$	Total
Success	$x_1$	$x_2$	...	$x_c$	$x$
Failure	$n_1 - x_1$	$n_2 - x_2$	...	$n_c - x_c$	$n - x$
Total	$n_1$	$n_2$	...	$n_c$	$n$

Common (average) estimate:

Under the null hypothesis, the estimate of  $p$  is:

$$\hat{p} = \frac{x}{n}$$



## Hypothesis test for several proportions

Common (average) estimate:

Under the null hypothesis the estimate of  $p$  is:

$$\hat{p} = \frac{x}{n}$$

"Use" this common estimate in each group:

If the null hypothesis is true, we expect that the  $j$ 'th group/sample has  $e_{1j}$  successes and  $e_{2j}$  failures, where

$$e_{1j} = n_j \cdot \hat{p} = \frac{n_j \cdot x}{n}$$

$$e_{2j} = n_j(1 - \hat{p}) = \frac{n_j \cdot (n - x)}{n}$$

## Hypothesis test for several proportions

Make a table with the *expected* counts for the  $c$  samples:

$e_{ij}$	Sample 1	Sample 2	...	Sample $c$	Total
Success	$e_{11}$	$e_{12}$	...	$e_{1c}$	$x$
Failure	$e_{21}$	$e_{22}$	...	$e_{2c}$	$n - x$
Total	$n_1$	$n_2$	...	$n_c$	$n$

General way to find the expected counts in frequency tables:

$$e_{ij} = \frac{(i\text{'th row total}) \cdot (j\text{'th column total})}{\text{total}}$$

## Computation of the test statistic - Method 7.20

The test statistic becomes

$$\chi_{\text{obs}}^2 = \sum_{i=1}^2 \sum_{j=1}^c \frac{(o_{ij} - e_{ij})^2}{e_{ij}}$$

where  $o_{ij}$  is the observed frequency in cell  $(i, j)$  and  $e_{ij}$  is the expected frequency in cell  $(i, j)$ .

## Find the $p$ -value or use the critical value - Method 7.20

Sampling distribution for test statistic under  $H_0$ :

$\chi^2$ -distribution with  $(c - 1)$  degrees of freedom (approx.)

Critical value method:

If  $\chi_{\text{obs}}^2 > \chi_{\alpha}^2(c - 1)$  the null hypothesis is rejected.

Rule of thumb for validity of the test:

All expected values  $e_{ij} \geq 5$ .

## Example 2 - continued

The *observed* values  $o_{ij}$

Observed	Blood clot	No Blood clot
B. C. pill	23	34
No B. C. pill	35	132

## Example 2 - continued

Use “the rule” for expected values four times, e.g.:

$$e_{22} = \frac{167 \cdot 166}{224} = 123.76$$

The *expected* values  $e_{ij}$ :

Expected	Blood clot	No Blood clot	Total
B. C. pill	14.76	42.24	57
No B. C. pill	43.24	123.76	167
Total	58	166	224

## Example 2 - continued

Compute the *expected* values  $e_{ij}$

Expected	Blood clot	No Blood clot	Total
B. C. pill			57
No B. C. pill			167
Total	58	166	224

## Example 2 - continued

The test statistic (remember to include all cells):

$$\chi_{\text{obs}}^2 = \frac{(23 - 14.76)^2}{14.76} + \frac{(34 - 42.24)^2}{42.24} + \frac{(35 - 43.24)^2}{43.24} + \frac{(132 - 123.76)^2}{123.76}$$

$$= 8.33$$

Critical value:

```
qchisq(0.95, 1)
```

```
[1] 3.841
```

Conclusion:

We reject the null hypothesis - there *is* a significantly higher risk of blood clots in the birth control pill group.

## Example 2 - continued

chisq.test for equality of two proportions in R

```
# Test whether probabilities are equal for the two groups
chisq.test(pill.study, correct = FALSE)

##
## Pearson's Chi-squared test
##
## data: pill.study
## X-squared = 8.3, df = 1, p-value = 0.004

# Expected values
chisq.test(pill.study, correct = FALSE)$expected

##           Blood Clot No Clot
## Pill           14.76  42.24
## No pill          43.24 123.76
```

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## Example 3: Analysis of contingency tables

A  $3 \times 3$  table - 3 samples, 3-category outcomes

	4 weeks bef	2 weeks bef	1 week bef
Candidate I	79	91	93
Candidate II	84	66	60
Undecided	37	43	47
	$n_1 = 200$	$n_2 = 200$	$n_3 = 200$

Are the votes equally distributed?

$$H_0: p_{i1} = p_{i2} = p_{i3}, \quad i = 1, 2, 3.$$

## Analysis of contingency tables

A  $3 \times 3$  table - 1 sample, two 3-category variables:

	bad	average	good
bad	23	60	29
average	28	79	60
good	9	49	63

Is there independence between the row and column variables?

$$H_0: p_{ij} = p_i \cdot p_j$$

## Computation of the test statistic – no matter the type of table 7.22

In a contingency table with  $r$  rows and  $c$  columns, the test statistic is:

$$\chi_{\text{obs}}^2 = \sum_{i=1}^r \sum_{j=1}^c \frac{(o_{ij} - e_{ij})^2}{e_{ij}}$$

where  $o_{ij}$  is the observed value in cell  $(i, j)$  and  $e_{ij}$  is the expected value in cell  $(i, j)$ .

General way to find the expected counts in frequency tables:

$$e_{ij} = \frac{(i\text{'th row total}) \cdot (j\text{'th column total})}{\text{total}}$$

## Find $p$ -value or use critical value - Method 7.22

Sampling distribution for test-statistic under  $H_0$ :

$\chi^2$ -distribution with  $(r-1)(c-1)$  degrees of freedom (approx.).

Critical value method:

If  $\chi_{\text{obs}}^2 > \chi_{\alpha}^2$  with  $(r-1)(c-1)$  degrees of freedom, the null hypothesis is rejected.

Rule of thumb for validity of the test:

All expected values  $e_{ij} \geq 5$ .

## Example 3 - continued

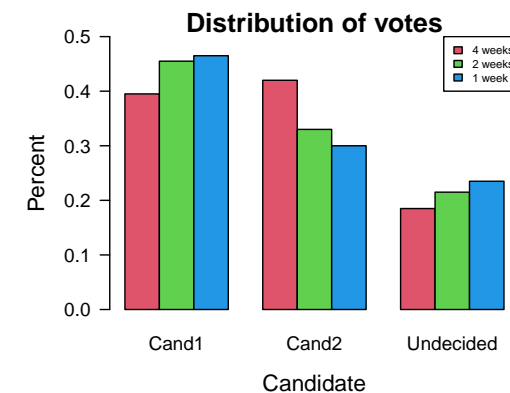
`chisq.test` for contingency tables

```
# Read data table into R
poll <-matrix(c(79, 91, 93, 84, 66, 60, 37, 43, 47),
             ncol = 3, byrow = TRUE)
colnames(poll) <- c("4 weeks", "2 weeks", "1 week")
rownames(poll) <- c("Cand1", "Cand2", "Undecided")

# Show column percentages
prop.table(poll, 2)

##           4 weeks 2 weeks 1 week
## Cand1      0.395  0.455  0.465
## Cand2      0.420  0.330  0.300
## Undecided  0.185  0.215  0.235
```

## Example 3 - continued



## Example 3 - continued

```
# Testing for same distribution in the three populations
chisq.test(poll, correct = FALSE)

##
## Pearson's Chi-squared test
##
## data:  poll
## X-squared = 7, df = 4, p-value = 0.1

# Expected values
chisq.test(poll, correct = FALSE)$expected

##           4 weeks 2 weeks 1 week
## Cand1      87.67   87.67  87.67
## Cand2      70.00   70.00  70.00
## Undecided  42.33   42.33  42.33
```

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