# 02417: Time Series Analysis Week 8 – the MARIMA estimation method and Multivariate time series

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#### DTU Compute

Based on material previous material from the course

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### Week 8: Outline of the lecture

- Introduction to MARIMA
- How the MARIMA method works (how it includes an MA part)
- Chapter 9 Multivariate time series

### The MARIMA package and the Spliid method

Run examples to find out how the marima and the Spliid method works!

- marima is an R package implementing the Spliid method, see: Article
- Let's run some examples together, download "Week8\_example\_marima.zip" and unzip

### The MARIMA package and the Spliid method

Remember ARMA is noise through a transfer function:



How the Spliid method includes the MA part:

Step 1: Make AR and estimate with LS

Step 2: Take the residuals and lag as MA part; Put in model with the AR part, and estimate again with LS

Step 3: Iterate until the residuals don't change

Start with "marima\_from\_scratch\_arma.R": Simulate ARMA and estimate

### The MARIMA package and the Spliid method

Run examples to find out how the marima and the Spliid method works!

- Can we use the Spliid method with an external regressor!?
- Yes, just include it in the LS model (with lags)
- We can fit an ARMAX model with a transfer function!
- ► E.g. ARMAX(1,1):

$$Y_t = -\phi_1 Y_{t-1} + \omega_1 x_{t-1} + \theta_1 \varepsilon_{t-1} + \varepsilon_t$$
$$Y_t = \frac{\omega_1 B}{1 + \phi_1 B} x_t + \frac{1 + \theta_1 B}{1 + \phi_1 B} \varepsilon_t$$

Look into "marima\_from\_scratch\_armax.R"

#### Multivariate models

Re-consider the univariate transfer function model:



$$Y_t = h(B)X_t + N_t$$

▶ What if there is a feedback from *Y* to *X*?

### Closed Loop Models



 $X_t = h_2(B) Y_t + N_{2,t}$ 

#### **Closed Loop Models**

$$Y_t = h_1(B)X_t + N_{1,t}$$
  
 $X_t = h_2(B)Y_t + N_{2,t}$ 

Or:

$$\begin{pmatrix} 1 & -h_1(B) \\ -h_2(B) & 1 \end{pmatrix} \begin{pmatrix} Y_t \\ X_t \end{pmatrix} = \begin{pmatrix} N_{1,t} \\ N_{2,t} \end{pmatrix}$$

Two inputs  $(N_1, N_2)$ ;

- Two outputs (Y, X);
- ► Four transfer functions from input to output.

## Predetor-pray: Mink-Muskrat example



Transfer from  $N_1$ ,  $N_2$  to Y:

$$Y_t = h_1(B)X_t + N_{1,t}$$
  

$$X_t = h_2(B)Y_t + N_{2,t}$$
  

$$Y_t = h_1(B)(h_2(B)Y_t + N_{2,t}) + N_{1,t}$$

Z-domain:

$$Y(z) = H_1(z)(N_2(z) + H_2(z)Y(z)) + N_1(z)$$
  

$$Y(z) - H_1(z)H_2(z)Y(z) = H_1(z)N_2(z) + N_1(z)$$
  

$$Y(z) = \frac{1}{1 - H_1(z)H_2(z)}N_1(z) + \frac{H_1(z)}{1 - H_1(z)H_2(z)}N_2(z)$$

Transfer functions from  $N_1$ ,  $N_2$  to Y:

$$N_1: rac{1}{1-H_1(z)H_2(z)}$$
 and  $N_2: rac{H_1(z)}{1-H_1(z)H_2(z)}$ 

#### Multivariate transfer function

Model equation:

$$\begin{pmatrix} 1 & -h_1(B) \\ -h_2(B) & 1 \end{pmatrix} \begin{pmatrix} Y_t \\ X_t \end{pmatrix} = \begin{pmatrix} N_{1,t} \\ N_{2,t} \end{pmatrix}$$

Model equation in Z-domain:

$$\begin{pmatrix} 1 & -H_1(z) \\ -H_2(z) & 1 \end{pmatrix} \begin{pmatrix} Y(z) \\ X(z) \end{pmatrix} = \begin{pmatrix} N_1(z) \\ N_2(z) \end{pmatrix}$$

Thus, from the result on the previous slide:

$$\begin{pmatrix} Y(z) \\ X(z) \end{pmatrix} = \frac{1}{1 - H_1(z)H_2(z)} \begin{pmatrix} 1 & H_1(z) \\ H_2(z) & 1 \end{pmatrix} \begin{pmatrix} N_1(z) \\ N_2(z) \end{pmatrix}$$

Multivariate transfer function:

$$\begin{pmatrix} \frac{1}{1-H_1(z)H_2(z)} & \frac{H_1(z)}{1-H_1(z)H_2(z)} \\ \frac{H_2(z)}{1-H_1(z)H_2(z)} & \frac{1}{1-H_1(z)H_2(z)} \end{pmatrix}$$

#### Multivariate ARMA models

The multivariate ARMA process

 $\boldsymbol{Y}_t + \boldsymbol{\phi}_1 \boldsymbol{Y}_{t-1} + \ldots + \boldsymbol{\phi}_p \boldsymbol{Y}_{t-p} = \boldsymbol{\epsilon}_t + \boldsymbol{\theta}_1 \boldsymbol{\epsilon}_{t-1} + \ldots + \boldsymbol{\theta}_q \boldsymbol{\epsilon}_{t-q}$ 

where  $\{\epsilon_t\}$  is white noise, is called a Vector ARMA (VARMA) process.

The model can be written

$$\boldsymbol{\phi}(B)(\boldsymbol{Y}_t - \boldsymbol{c}) = \boldsymbol{\theta}(B)\boldsymbol{\epsilon}_t$$

- The individual time series may have been transformed and differenced
- The variance-covariance matrix of the multivariate white noise process  $\{\epsilon_t\}$  is denoted  $\Sigma$ .
- The matrices  $\phi(B)$  and  $\theta(B)$  have elements which are polynomials in the backshift operator
- The diagonal elements have leading terms of unity
- ▶ The off-diagonal elements have leading terms of zero (i.e. they normally start in B)

### Air pollution in cities NO and $NO_2$

$$\begin{bmatrix} X_{1,t} \\ X_{2,t} \end{bmatrix} = \begin{bmatrix} 0.9 & -0.1 \\ 0.4 & 0.8 \end{bmatrix} \begin{bmatrix} X_{1,t-1} \\ X_{2,t-1} \end{bmatrix} + \begin{bmatrix} \xi_{1,t} \\ \xi_{2,t} \end{bmatrix}, \quad \mathbf{\Sigma} = \begin{bmatrix} 30 & 21 \\ 21 & 23 \end{bmatrix}$$

Matrix formulation using the backshift operator:

$$\begin{bmatrix} 1 - 0.9B & 0.1B \\ -0.4B & 1 - 0.8B \end{bmatrix} \boldsymbol{X}_t = \boldsymbol{\xi}_t \quad \text{or} \quad \boldsymbol{\phi}(B) \boldsymbol{X}_t = \boldsymbol{\xi}_t$$

### Stationarity and Invertability

The multivariate ARMA process

$$\boldsymbol{\phi}(B)(\boldsymbol{Y}_t - \boldsymbol{c}) = \boldsymbol{\theta}(B)\boldsymbol{\epsilon}_t$$

is stationary if

$$\det(\boldsymbol{\phi}(z^{-1})) = 0 \quad \Rightarrow \quad |z| < 1$$

is invertible if

$$\det(\boldsymbol{\theta}(z^{-1})) = 0 \quad \Rightarrow \quad |z| < 1$$

#### Auto Covariance Matrix Functions

$$\boldsymbol{\Gamma}_{k} = E[(\boldsymbol{Y}_{t-k} - \boldsymbol{\mu}_{Y})(\boldsymbol{Y}_{t} - \boldsymbol{\mu}_{Y})^{T}] = \boldsymbol{\Gamma}_{-k}^{T}$$

Example for bivariate case  $\boldsymbol{Y}_t = (Y_{1,t} \ Y_{2,t})^T$ :

$$\boldsymbol{\mathsf{\Gamma}}_{k} = \left[ \begin{array}{cc} \gamma_{11}(k) & \gamma_{12}(k) \\ \gamma_{21}(k) & \gamma_{22}(k) \end{array} \right] = \left[ \begin{array}{cc} \gamma_{11}(k) & \gamma_{12}(k) \\ \gamma_{12}(-k) & \gamma_{22}(k) \end{array} \right]$$

We can describe these by plotting

- each autocovariance or autocorrelation function for k = 0, 1, 2, ... and
- each cross-covariance or cross-correlation function for  $k = 0, \pm 1, \pm 2, \ldots$

### Identification using Autocovariance Matrix Functions

- Sample Correlation Matrix Function;  $R_k$  near zero for pure moving average processes of order q when k > q
- Sample Partial Correlation Matrix Function;  $S_k$  near zero for pure autoregressive processes of order p when k > p

## Identification using (multivariate) prewhitening

- Fit univariate models to each individual series
- Investigate the residuals as a multivariate time series
- Model selection procedure!

The cross correlations can then be compared with  $\pm 2/\sqrt{N}$ This is **not** the same form of prewhitening as in Chapter 8 Remember the result in two dimensions:

$$Y_t = h_1(B)X_t + N_{1,t}$$
(1)

$$X_t = h_2(B) Y_t + N_{2,t}$$
(2)

$$Y_t = h_1(B)(h_2(B)Y_t + N_{2,t}) + N_{1,t}$$
(3)

in general the multivariate model  $\boldsymbol{\phi}(B) Y_t = \boldsymbol{\theta}(B) \epsilon_t$  is equivalent to

$$\operatorname{diag}(\operatorname{det}(\boldsymbol{\phi}(B)))Y_t = \operatorname{adj}(\boldsymbol{\phi}(B))\boldsymbol{\theta}(B)\boldsymbol{\epsilon}_t$$

Therefore the corresponding univariate models will have much higher order, so although this is often done in the literature: Don't take this approach!

### Multivariate ARMA(p,q) processes (centered data)

▶ Matrices with polynomials in *B* as elements:

$$\boldsymbol{\phi}(B) \boldsymbol{Y}_t = \boldsymbol{\theta}(B) \boldsymbol{\epsilon}_t$$

So the coefficients are now matrices:

$$\boldsymbol{Y}_t + \boldsymbol{\phi}_1 \boldsymbol{Y}_{t-1} + \ldots + \boldsymbol{\phi}_p \boldsymbol{Y}_{t-p} = \boldsymbol{\epsilon}_t + \boldsymbol{\theta}_1 \boldsymbol{\epsilon}_{t-1} + \ldots + \boldsymbol{\theta}_q \boldsymbol{\epsilon}_{t-q}$$

- In general, no analytic solution exits.
- Therefore, estimation algorithms ( or numerical optimization) is necessary.

#### Estimation procedures

For multivariate ARX(p):

Least squares estimation is possible

For multivariate ARMAX(p,q):

- The Spliid method (Henrik Spliid, 1983)
- Maximum likelihood

Go and have a look into marima\_from\_scratch\_armax\_bivariate.R, it's a very short example on how to simulate and fit a bivariate ARMAX(1,1) with MARIMA.

### Highlights

Closed loop model as multivariate transfer function

$$\begin{pmatrix} 1 & -h_1(B) \\ -h_2(B) & 1 \end{pmatrix} \begin{pmatrix} Y_t \\ X_t \end{pmatrix} = \begin{pmatrix} N_{1,t} \\ N_{2,t} \end{pmatrix}$$

Multivariate ARMA models

$$\boldsymbol{\phi}(B)(\boldsymbol{Y}_t - \boldsymbol{c}) = \boldsymbol{\theta}(B)\boldsymbol{\epsilon}_t$$

is stationary if

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Auto covariance matrix functions

$$\boldsymbol{\Gamma}_{k} = E[(\boldsymbol{Y}_{t-k} - \boldsymbol{\mu}_{Y})(\boldsymbol{Y}_{t} - \boldsymbol{\mu}_{Y})^{T}] = \boldsymbol{\Gamma}_{-k}^{T}$$

All VARMA models can be written as VAR(1)

#### Exercises and Assignment 3

- A new exercise was uploaded, gives you "hands-on" of fitting ARX and ARMAX to data from an experimental set up
- Upload Assignment 3 in the afternoon, take a look at it