
The non-central Nakagami distribution

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Abstract

The Nakagami distribution describe the square root of a random variable drawn from a Gamma distribution. Equivalently, the Gamma distribution can be seen as a one-dimensional Wishart distribution. In this note, we consider the distribution of the square root of a random variable drawn from a non-central one-dimensional Wishart distribution. We present the probability density function of this distribution along with closed-form expressions for its moments.

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1 The Nakagami distribution

Let $z_d \sim \mathcal{N}(0, \sigma^2)$, $d = 1, \dots, D$ be iid samples from a zero-mean normal distribution. Then

$$x = \sum_{d=1}^D z_d^2 \in \mathbb{R}^+ \quad (1.1)$$

follows a Gamma distribution. Equivalently, and more suitable for our purposes, x also follows a one-dimensional Wishart distribution [1],

$$x \sim \mathcal{W}_1(D, \sigma^2) \quad (1.2)$$

$$p(x) = \mathcal{W}_1(x | D, \sigma^2) = \sqrt{\frac{x^{D-2}}{2^D \sigma^{2D}}} \Gamma(D/2)^{-1} \exp\left(-\frac{x}{2\sigma^2}\right). \quad (1.3)$$

This distribution can be seen as describing the squared norm of a zero-mean normally distributed vector. The norm of this vector is then given by $y = \sqrt{x}$, which follows a Nakagami distribution [2],

$$y \sim \text{Nakagami}(D/2, D\sigma^2) \quad (1.4)$$

$$p(y) = \text{Nakagami}(y | D/2, D\sigma^2) = \frac{2}{\Gamma(D/2) (\sqrt{2}\sigma)^D} y^{D-1} \exp\left(-\frac{y^2}{2\sigma^2}\right). \quad (1.5)$$

This expression is easily derived by the change of value theorem; see next section.

The expectation and variance of a Nakagami distributed variable is

$$\mathbb{E}[y] = \frac{\Gamma\left(\frac{D+1}{2}\right)}{\Gamma\left(\frac{D}{2}\right)} \sqrt{2}\sigma \quad (1.6)$$

$$\text{var}[y] = \sigma^2 \left(D - 2 \left(\frac{\Gamma\left(\frac{(D+1)}{2}\right)}{\Gamma\left(\frac{D}{2}\right)} \right)^2 \right). \quad (1.7)$$

We will derive a more general version of these results in the next section.

2 The non-central Nakagami distribution

Now, let $z_d \sim \mathcal{N}(\mu_d, \sigma^2)$, $d = 1, \dots, D$ be independent samples from normal distributions with variance σ^2 . Then

$$x = \sum_{d=1}^D z_d^2 \in \mathbb{R}^+ \quad (2.1)$$

follows a one-dimensional non-central Wishart distribution [1],

$$x \sim \mathcal{W}_1(D, \sigma^2, \Omega) \quad (2.2)$$

$$\Omega = \frac{\sum_{d=1}^D \mu_d^2}{\sigma^2} \quad (2.3)$$

$$\begin{aligned} p(x) &= \mathcal{W}_1(x | D, \sigma^2, \Omega) \\ &= \sqrt{\frac{x^{D-2}}{2^D \sigma^{2D}}} \Gamma(D/2)^{-1} {}_0F_1(D/2, 1/4\Omega\sigma^{-2}x) \exp\left(-\frac{x}{2\sigma^2}\right) \exp\left(-\frac{\Omega}{2}\right). \end{aligned} \quad (2.4)$$

Here ${}_0F_1$ is a generalized hypergeometric function; see Sec. 7.3 of Muirhead's book [1]. This distribution can be seen as describing the squared norm of a normal distributed vector with isotropic variance. The norm of this vector is then given by $y = \sqrt{x}$, which we say follows a non-central Nakagami distribution. This distribution can be derived by change-of-variables to be

$$p(y) = \mathcal{W}_1(y^2 | D, \sigma^2, \Omega) \cdot 2y \quad (2.5)$$

$$= \frac{y^{D-1}}{2^{(D-2)/2} \sigma^D \Gamma(D/2)} {}_0F_1\left(\frac{D}{2}, \Omega \frac{y^2}{4\sigma^2}\right) \exp\left(-\frac{y^2}{2\sigma^2}\right) \exp\left(-\frac{\Omega}{2}\right). \quad (2.6)$$

To compute the moments of \sqrt{x} we recall the following result from Muirhead [1, Theorem 10.3.7]. Let $\mathbf{X} \in \mathbb{R}^{q \times q}$ and $\mathbf{X} \sim \mathcal{W}_q(D, \Sigma, \Omega)$ then

$$\mathbb{E}[(\det(\mathbf{X}))^k] = (\det(\Sigma))^k 2^{qk} \frac{\Gamma_q(D/2 + k)}{\Gamma_q(D/2)} {}_1F_1(-k, D/2, -1/2\Omega), \quad (2.7)$$

where ${}_1F_1$ is a generalized hypergeometric function. Since $y = \sqrt{x}$ is a positive scalar, its determinant is merely y , and we get

$$\mathbb{E}[\sqrt{x}^{2k}] = \sigma^{2k} 2^k \frac{\Gamma(D/2 + k)}{\Gamma(D/2)} {}_1F_1(-k, D/2, -1/2\Omega). \quad (2.8)$$

From this we see that the mean and the variance of $y = \sqrt{x}$ is

$$\mathbb{E}[\sqrt{x}] = \sigma 2^{1/2} \frac{\Gamma((D+1)/2)}{\Gamma(D/2)} {}_1F_1(-1/2, D/2, -1/2\Omega) \quad (2.9)$$

$$\mathbb{E}[x] = \sigma^2 2 \frac{\Gamma((D+2)/2)}{\Gamma(D/2)} {}_1F_1(-1, D/2, -1/2\Omega) \quad (2.10)$$

$$\text{var}[\sqrt{x}] = \mathbb{E}[x] - \mathbb{E}[\sqrt{x}]^2 \quad (2.11)$$

$$\begin{aligned} &= \frac{\sigma^2 2}{\Gamma(D/2)} \left(\Gamma\left(\frac{D+2}{2}\right) {}_1F_1(-1, D/2, -1/2\Omega) \right. \\ &\quad \left. - \frac{\Gamma((D+1)/2)^2}{\Gamma(D/2)} {}_1F_1(-1/2, D/2, -1/2\Omega)^2 \right). \end{aligned} \quad (2.12)$$

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References

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